

PERSUASION WITH CORRELATION NEGLECT: A FULL MANIPULATION
RESULT

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ONLINE APPENDIX

A1. Proof of Lemma 2

A receiver with correlation neglect believes that the conditional joint distribution of the signals is the product of the conditional marginals:

$$q(s_1, s_2, \dots, s_m | \omega) = \prod_{i=1}^m q_i(s_i | \omega).$$

Therefore upon observing realisation $s = (s_1, \dots, s_m)$ which leads to posteriors $\boldsymbol{\mu} = (\mu_1, \dots, \mu_m)$ with $\mu_i(\omega) = (p(\omega)q_i(s_i | \omega)) / (\sum_{\nu} p(\nu)q_i(s_i | \nu))$, her posterior belief is:

$$\begin{aligned} \frac{p(\omega)q(s_1, \dots, s_m | \omega)}{\sum_{\nu} p(\nu)q(s_1, \dots, s_m | \nu)} &= \frac{p(\omega) \prod_{i=1}^m q_i(s_i | \omega)}{\sum_{\nu} p(\nu) \prod_{i=1}^m q_i(s_i | \nu)} \\ &= \frac{(1/p(\omega)^{m-1}) \prod_{i=1}^m p(\omega)q_i(s_i | \omega) / \sum_{\nu} p(\nu)q_i(s_i | \nu)}{\sum_{\nu} (1/p(\nu)^{m-1}) \prod_{i=1}^m p(\nu)q_i(s_i | \nu) / \sum_{\nu} p(\nu)q_i(s_i | \nu)} \\ &= \frac{\prod_{i=1}^m \mu_i(\omega) / p(\omega)^{m-1}}{\sum_{\nu \in \Omega} \prod_{i=1}^m \mu_i(\nu) / p(\nu)^{m-1}} \end{aligned}$$

Hence we can write,

$$\mu^{CN}(\boldsymbol{\mu})(\omega) = \frac{\prod_{i=1}^m \mu_i(\omega) / p(\omega)^{m-1}}{\sum_{\nu \in \Omega} \left(\prod_{i=1}^m \mu_i(\nu) / p(\nu)^{m-1} \right)}$$

□

A2. Additional proofs for Theorem 1

We prove that the joint conditional distribution described in (10) is well defined.

First, we show that $0 \leq \gamma^\omega \leq 1$. From the definition, γ^ω is trivially positive. Moreover, it cannot be that both $\alpha/z_\omega^\omega > 1$ and $\beta/(\max_{\nu \neq \omega} \{z_\nu^\omega\}) > 1$. If that were the case, then $z_\omega^\omega < \alpha$ and $\max_{\nu \neq \omega} \{z_\nu^\omega\} < \beta$. But then $\sum_{\nu} z_\nu^\omega \leq z_\omega^\omega + (n-1) \max_{\nu \neq \omega} \{z_\nu^\omega\} < \alpha + (n-1)\beta = 1$, which contradicts that $\sum_{\nu} z_\nu^\omega = 1$. Therefore, $\gamma^\omega \leq 1$.

Second we show that both $0 \leq \lambda_\omega^\omega \leq 1$ and $0 \leq \lambda_\nu^\omega \leq 1$. Again, from the definition it is trivial to see that these numbers are below 1. To see that they are positive note that, $\lambda_\omega^\omega = \alpha - \gamma^\omega z_\omega^\omega \geq \alpha - (\alpha/z_\omega^\omega)z_\omega^\omega = 0$ and $\lambda_\nu^\omega = \beta - \gamma^\omega z_\nu^\omega \geq \beta - (\beta/z_\nu^\omega)z_\nu^\omega = 0$.

Lastly, $\gamma^\omega + \sum_v \lambda_v^\omega = \gamma^\omega(1 - \sum_v z_v^\omega) + \alpha + (n-1)\beta = 1$, therefore $\tau(\cdot | \omega)$ is indeed a distribution.

We now address the case in which ρ^ω is not interior. Consider a sequence of interior posteriors $\{\rho_m^\omega\}_{m \in \mathbb{N}}$ converging to ρ^ω and we replicate the proof by replacing in equation (6) ρ^ω by ρ_m^ω . Hence with probability converging to one, the information structure generates vectors of proportions $z^\omega(m)$ that satisfy:

$$0 \leq \lim_{m \rightarrow \infty} |\mu^{CN}(z^\omega(m)) - \rho^\omega| \leq \lim_{m \rightarrow \infty} |\mu^{CN}(z^\omega(m)) - \rho_m^\omega| + \lim_{m \rightarrow \infty} |\rho_m^\omega - \rho^\omega| = 0.$$

□

A3. Proof of Corollary 1

Recall from Section II.B that for each $\omega \in \Omega$, we defined $v^\omega(\mu) = \max_{a \in A_\mu} v(a, \omega)$ and that $v^\omega(\cdot)$ is continuous for all $\mu \in \Delta(\Omega)$ but for a finite set of posteriors. If ρ^ω is such that v^ω is continuous at ρ^ω , then given Theorem 1, the expected utility of the sender conditional on state ω converges to \bar{v}^ω .

Suppose now that ρ^ω is one of the finite points for which $A_{\rho^\omega}^\omega$ is not continuous. Then, by Assumption 1 there exists $a \in A_{\rho^\omega}^\omega$ and sequences $\{\rho_m^\omega\}_{m=1}^\infty, \{a_m\}_{m=1}^\infty$ with $\rho_m^\omega \neq \rho^\omega$ and $a_m \in A_{\rho_m^\omega}^\omega$, such that $\rho_m^\omega \rightarrow_{m \rightarrow \infty} \rho^\omega$ and $a_m \rightarrow_{m \rightarrow \infty} a$. Note that without loss of generality we can assume that for any m , v^ω is continuous at ρ_m^ω . Now, we use the proof of Theorem 1 in which at each m we replace ρ^ω in equation (6) by ρ_m^ω . Therefore, with probability converging to one, the information structure generates vectors of proportions $z^\omega(m)$ that satisfy:

$$|v(a_{\mu^{CN}(z^\omega(m))}, \omega) - \bar{v}^\omega| \leq |v(a_{\mu^{CN}(z^\omega(m))}, \omega) - v(a_m, \omega)| + |v(a_m, \omega) - \bar{v}^\omega| \rightarrow_{m \rightarrow \infty} 0.$$

Therefore, the ex-ante expected utility of the sender converges to $\sum_{\omega \in \Omega} p(\omega) \bar{v}^\omega$. □

A4. Proof of Proposition 1

Suppose first that $\rho(\omega) \neq 0$ for all $\omega \in \Omega$. Let $\hat{\mu}_m \in \Delta(\Omega)$ be defined in the following way:

$$\hat{\mu}_m(\omega) = \frac{\rho(\omega)^{\frac{1}{m}} p(\omega)^{\frac{m-1}{m}}}{\sum_{v \in \Omega} \rho(v)^{\frac{1}{m}} p(v)^{\frac{m-1}{m}}},$$

which then implies that:

$$\mu_m^{FC}(\hat{\mu}_m) = \rho.$$

Note that since p is interior, we can always design a signal structure τ^m of fully positively correlated signals that have two vectors of posteriors in the support, $\hat{\mu}_m = (\hat{\mu}_m, \dots, \hat{\mu}_m)$ and $\mu'_m = (\mu'_m, \dots, \mu'_m)$, where μ'_m is at a fixed distance $\delta > 0$ from p . The weights are then pinned down by the Bayesian Plausibility constraint such that:

$$\tau^m(\hat{\mu}_m) \hat{\mu}_m + (1 - \tau^m(\hat{\mu}_m)) \mu'_m = p.$$

Note that when $m \rightarrow \infty$, then $\hat{\mu}_m$ is arbitrarily close to p , as

$$\hat{\mu}_m(\omega) = \frac{p(\omega)^{1-\frac{1}{m}}}{\sum_{v \in \Omega} \left(\frac{\rho(v)}{\rho(\omega)}\right)^{\frac{1}{m}} p(v)^{1-\frac{1}{m}}} \xrightarrow{m \rightarrow \infty} p(\omega).$$

As a result, given the Bayesian Plausibility constraint, and maintaining μ'_m always at a fixed distance $\delta > 0$ away from p , $\tau^m(\hat{\mu}_m) \rightarrow_{m \rightarrow \infty} 1$. This implies that for any $\epsilon > 0$,

$$\lim_{m \rightarrow \infty} \tau^m(\{\mu = (\mu, \dots, \mu) \text{ s.t. } \tau^m(\mu) > 0 \mid |\mu_m^{FC}(\mu) - \rho| < \epsilon\}) = 1.$$

Lastly, consider the case in which $\rho(\omega) = 0$ for some $\omega \in \Omega$. Consider a sequence $\{\rho_m\}_{m=1}^{\infty}$ such that for any m , $\rho_m(v) > 0$ for any $v \in \Omega$ and $\lim_{m \rightarrow \infty} \rho_m = \rho$. Moreover, we can choose the sequence to satisfy for any $v, \omega \in \Omega$, $\lim_{m \rightarrow \infty} (\rho_m(v)/\rho_m(\omega))^{1/m} = 1$.

Let $\hat{\mu}_m \in \Delta(\Omega)$ be defined in the following way:

$$\hat{\mu}_m(\omega) = \frac{(\rho_m(\omega))^{\frac{1}{m}} p(\omega)^{\frac{m-1}{m}}}{\sum_{v \in \Omega} (\rho_m(v))^{\frac{1}{m}} p(v)^{\frac{m-1}{m}}},$$

which then implies that:

$$\mu_m^{FC}(\hat{\mu}_m) = \rho_m.$$

Note that since p is interior, we can always design a signal structure τ^m of fully positively correlated signals with two vectors of posteriors in the support, $\hat{\mu}_m = (\hat{\mu}_m, \dots, \hat{\mu}_m)$ and $\mu'_m = (\mu'_m, \dots, \mu'_m)$, where μ'_m is at a fixed distance $\delta > 0$ from p . The weights are then pinned down by the Bayesian Plausibility constraint such that:

$$\tau^m(\hat{\mu}_m)\hat{\mu}_m + (1 - \tau^m(\hat{\mu}_m))\mu'_m = p.$$

Note that when $m \rightarrow \infty$, then $\hat{\mu}_m$ is arbitrarily close to p , as

$$\hat{\mu}_m(\omega) = \frac{p(\omega)^{\frac{m-1}{m}}}{\sum_{v \in \Omega} \left(\frac{\rho_m(v)}{\rho_m(\omega)}\right)^{\frac{1}{m}} p(v)^{\frac{m-1}{m}}} \xrightarrow{m \rightarrow \infty} p(\omega).$$

As a result, given the Bayesian Plausibility constraint, and maintaining μ'_m always at a fixed distance $\delta > 0$ away from p , $\tau^m(\hat{\mu}_m) \rightarrow_{m \rightarrow \infty} 1$. As $\lim_{m \rightarrow \infty} \rho_m = \rho$, this implies that for any $\epsilon > 0$,

$$\lim_{m \rightarrow \infty} \tau^m(\{\mu = (\mu, \dots, \mu) \text{ s.t. } \tau^m(\mu) > 0 \mid |\mu_m^{FC}(\mu) - \rho| < \epsilon\}) = 1.$$

□