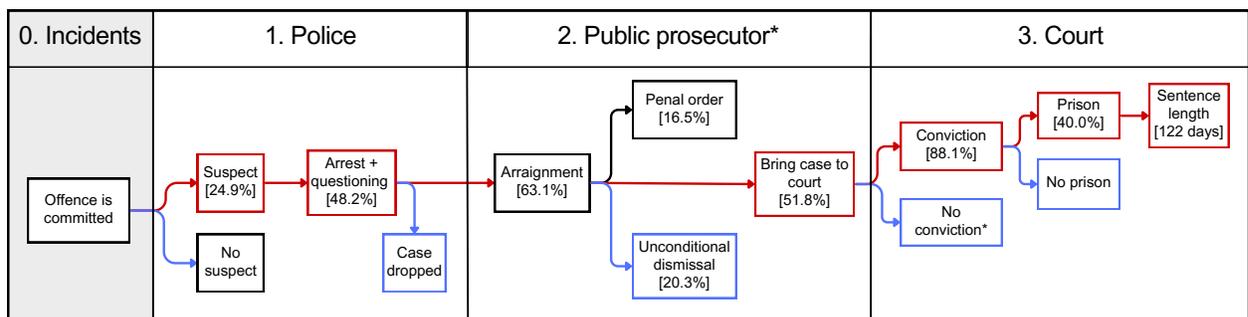


Supplemental Appendix to “Minority Salience and Criminal Justice Decisions”

Kyra Hanemaaijer, Nadine Ketel and Olivier Marie

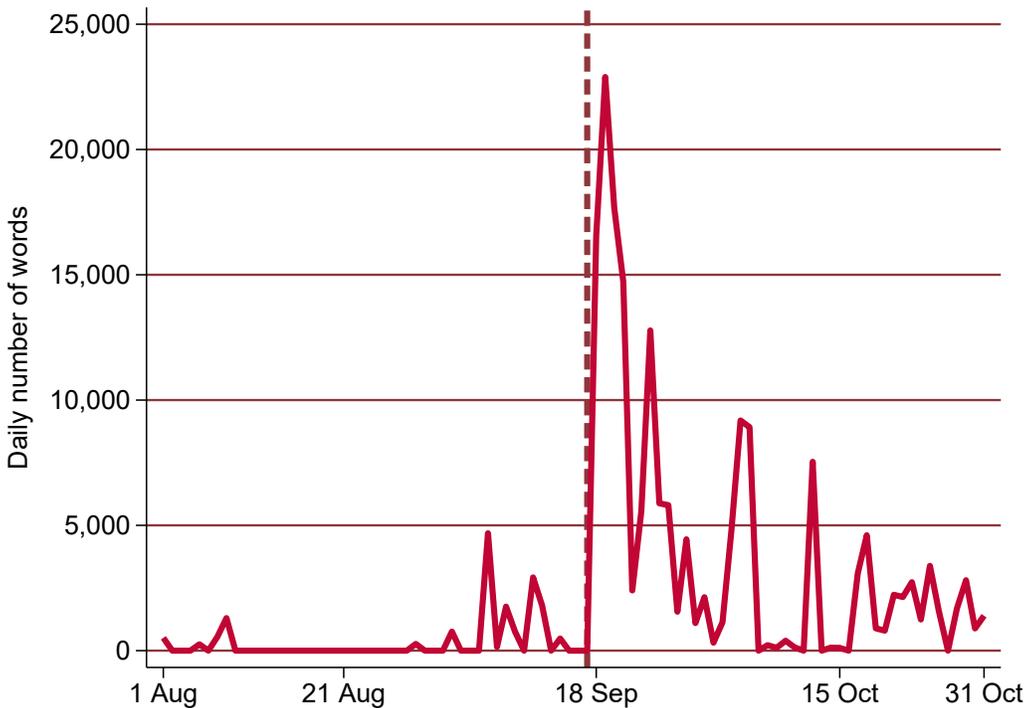
A Additional Figures

Figure A1: Flowchart of the criminal justice system in the Netherlands



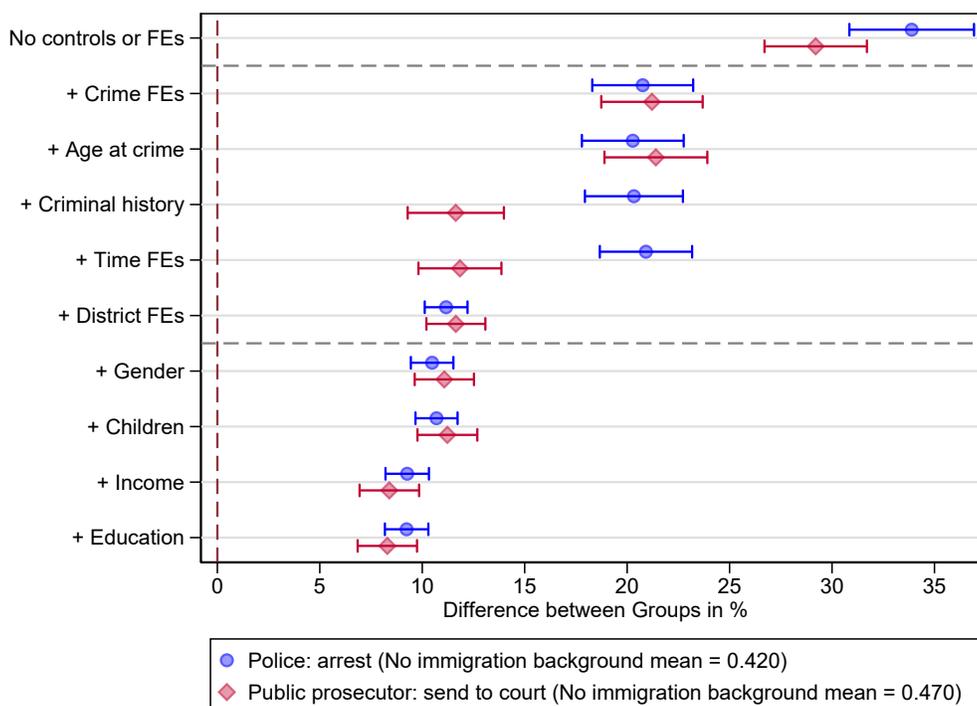
Notes: This figure shows the most important decisions (by stage) in the criminal justice system (CJS) in the Netherlands and the share of these decisions (in square brackets) conditional on the previous decision/stage. Less common decisions of the public prosecutor are not shown in the graph. These decision options are: conditional dismissal [5.1%], transaction [3.8%], other options [2.6%]. Reasons that judges do not convict include: the judge is not authorized, the public prosecutor is inadmissible, acquittal, dismissal from prosecution, guilty plea without punishment, and suspect inadmissible. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands on decisions in the CJS between 2014–2019, starting from the 5,619,477 incidents (crimes recorded by the police) in this period.

Figure A2: Press coverage linking crime and Moroccan identity in national Dutch newspapers



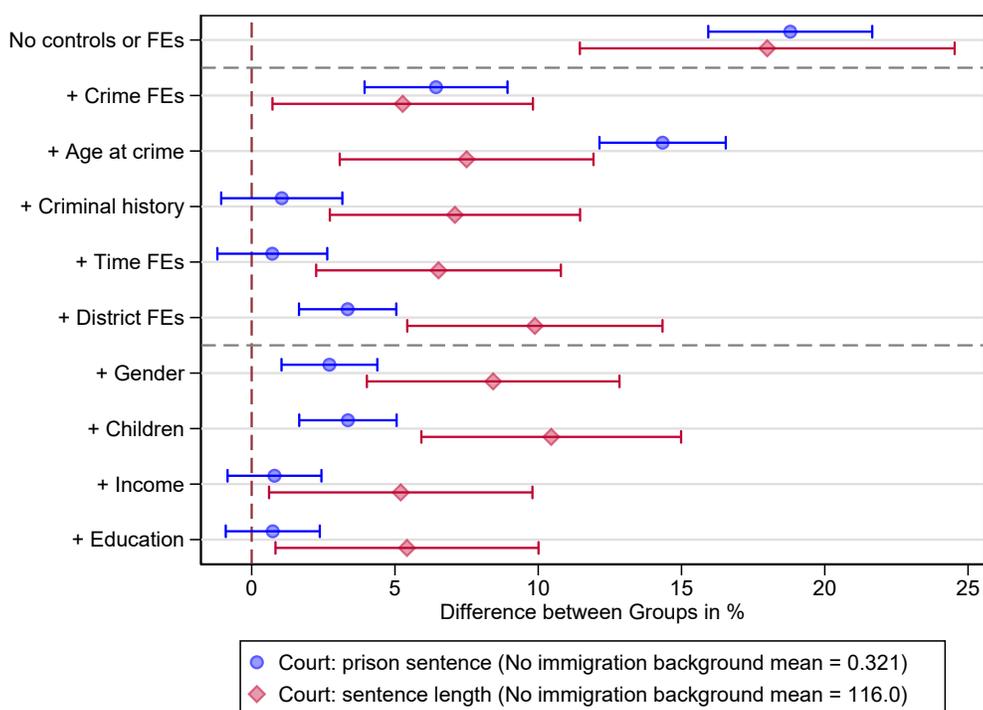
Notes: This figure plots the raw means of the daily number of words of news articles that contain words related to both “Moroccan” and “crime” in Dutch national newspapers between August 1 and October 31, 2019. The dashed vertical line marks the shock date (September 18). Source: Results based on calculations by the authors using Nexis Uni data. See Online Appendix for more details.

Figure A3: Disparities in police decisions to arrest suspect and public prosecutor decisions to send case to court between defendants of Moroccan descent and those without an immigration background, in detail



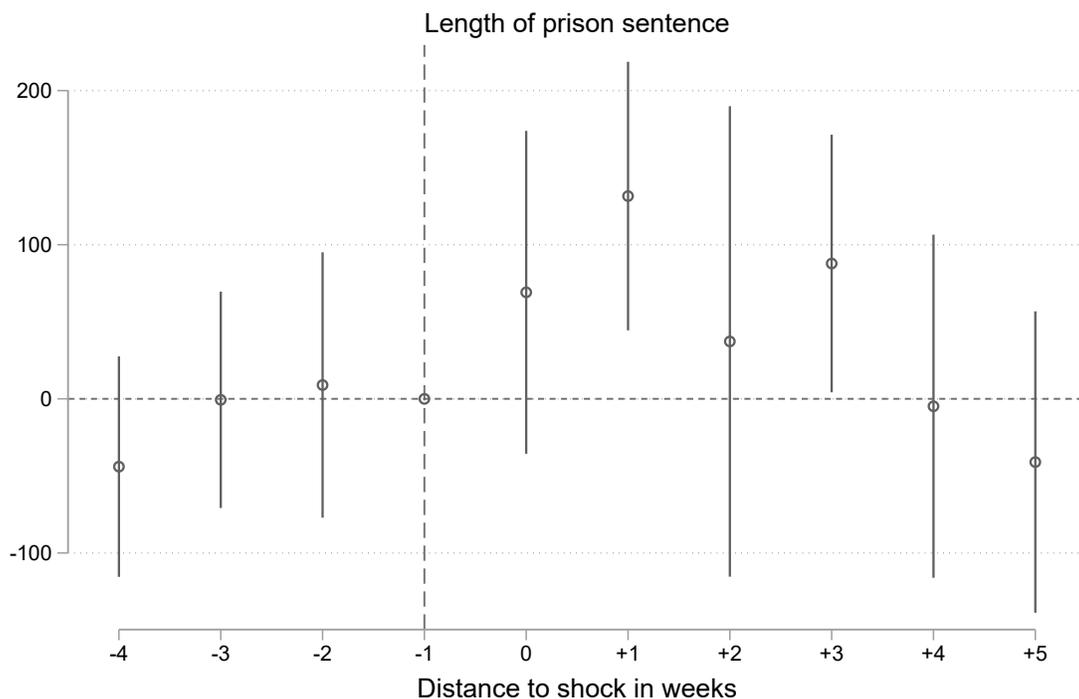
Notes: This figure plots the estimated coefficients and 95% confidence intervals for α_1 in the following equation: $Y_{i,c,t} = \alpha_1 Moroccan_i + \epsilon_{p,t}$ with $Y_{i,c,t}$ being the decision of the police to arrest a suspect i in case c at time t , or the decision of the public prosecutor to send a defendant i to court in case c at time t . α_1 represents the difference between defendants of Moroccan descent compared to those without an immigration background. Standard errors are clustered at the district-year level. Point estimates and CIs are transformed to differences in percentages relative to the mean outcome for suspects without an immigration background. Control variables and fixed effects are included sequentially. See Figure 1 and Data Appendix C for more information on exact control variables included. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands on decisions made between January 1, 2014 to September 17, 2019 (shock date, see Section 2.2).

Figure A4: Disparities in court decisions to impose prison sentence and sentence length between defendants of Moroccan descent and those without an immigration background, in detail



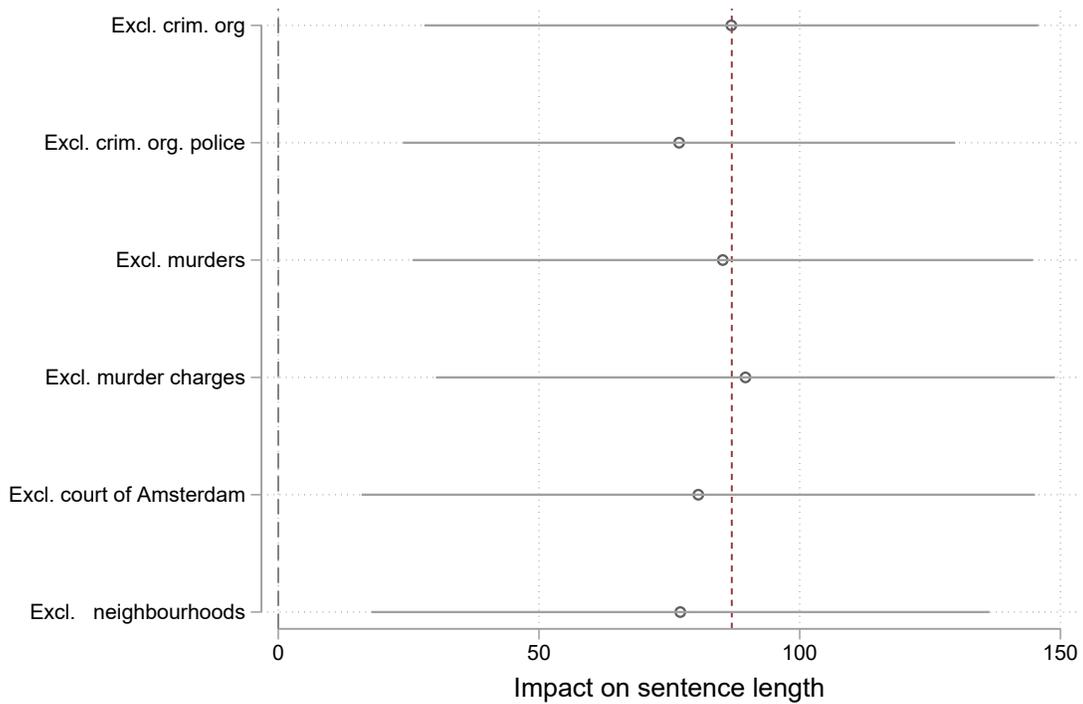
Notes: This figure plots the estimated coefficients and 95% confidence intervals for α_1 in the following equation: $Y_{i,c,t} = \alpha_1 Moroccan_i + \epsilon_{p,t}$ with $Y_{i,c,t}$ being the decision of a court to give a prison sentence and the decision on sentence length to defendant i in case c at time t . α_1 represents the difference between defendants of Moroccan descent compared to those without an immigration background. Standard errors are clustered at the district-year level. Point estimates and CIs are transformed to differences in percentages relative to the mean outcome for suspects without an immigration background. Control variables and fixed effects are included sequentially. See Figure 1 and Data Appendix C for more information on exact control variables included. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands on decisions made between January 1, 2014 to September 17, 2019 (shock date, see Section 2.2)

Figure A5: Impact of shock in salience on sentence length for descendants of Moroccan descent, by week



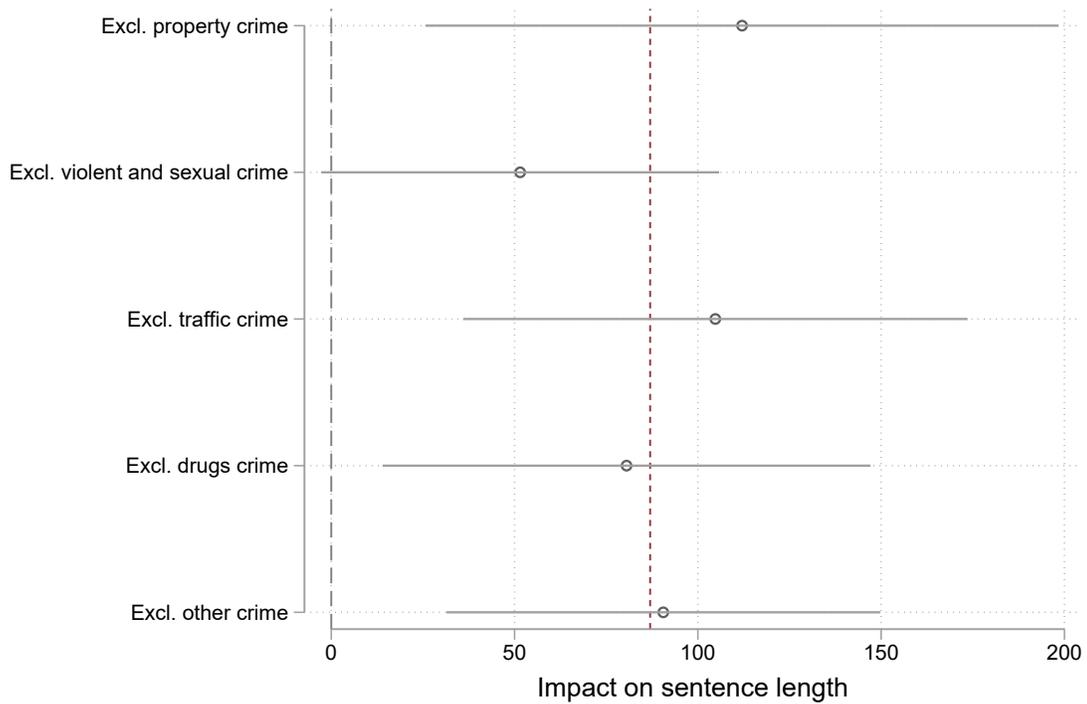
Notes: This figure plots the estimated coefficients $\theta_{i,j}$ and 95% confidence intervals (CIs) for the regression corresponding to the following equation: $Y_{i,c,t} = \alpha_1 Moroccan_i + \alpha_2 Post_t + \alpha_3 Moroccan_i \times Post_t + ShockYear_t[\beta_1 Moroccan_i + \sum_{j=-4, j \neq -1}^5 (\zeta_j T_j + \theta_{i,j} Moroccan_i \times T_j)] + X'_{i,c,t} \gamma + \delta_t + \delta_r + \delta_p + \epsilon_{p,t}$ with sentence length as the dependent variable. We use the sample conditional on prison sentence. The dashed vertical line marks the shock week. Robust standard errors clustered at the judicial district-year level are in parentheses. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A6: Impact of shock in salience on sentence length for defendants of Moroccan descent, using different specifications as robustness checks



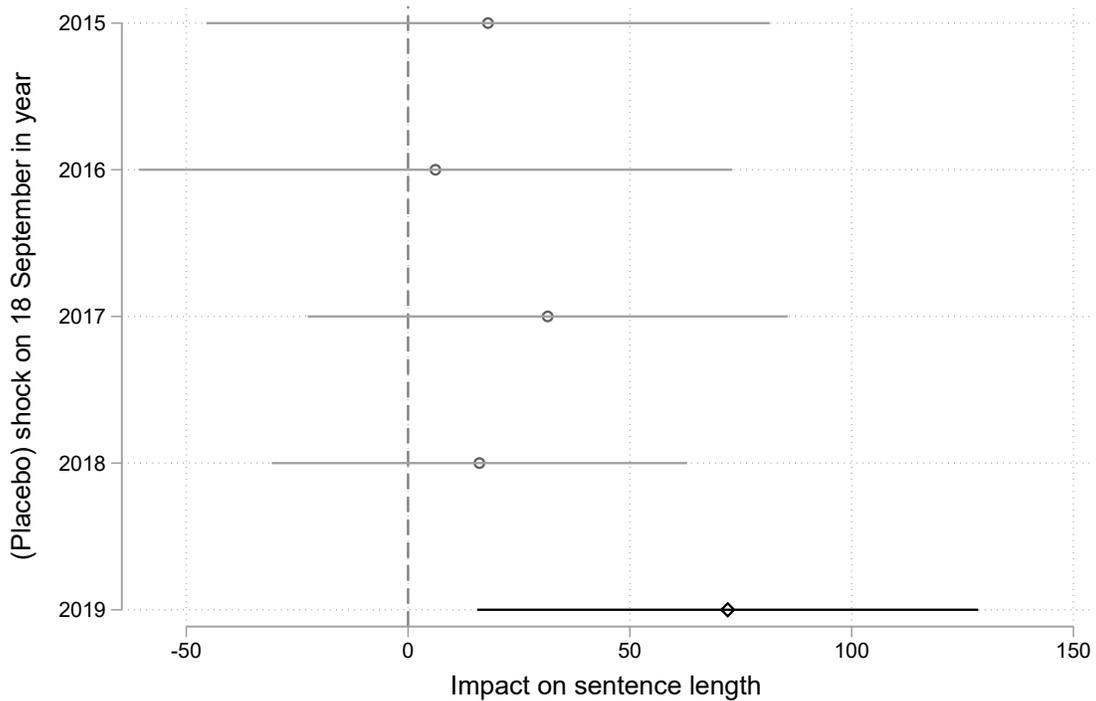
Notes: This figure plots the estimated coefficients β_3 and 95% confidence intervals for the regression corresponding to Equation 1 with sentence length (conditional on prison) as the dependent variable. We drop observations from our sample depending on the specification: excluding suspects involved in criminal organizations according to court data, excluding suspects involved in criminal organizations according to police data, excluding murders, excluding murders as charged by the public prosecutor, excluding cases at the court of Amsterdam, and excluding suspects living in neighbourhoods with links to the “Moroccan mafia”. The dashed red vertical line marks the effect size estimated with the main specification (in Table 3). Standard errors are clustered at the judicial district-year level. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A7: Impact of shock in salience on sentence length for defendants of Moroccan descent, excluding crime categories as robustness checks



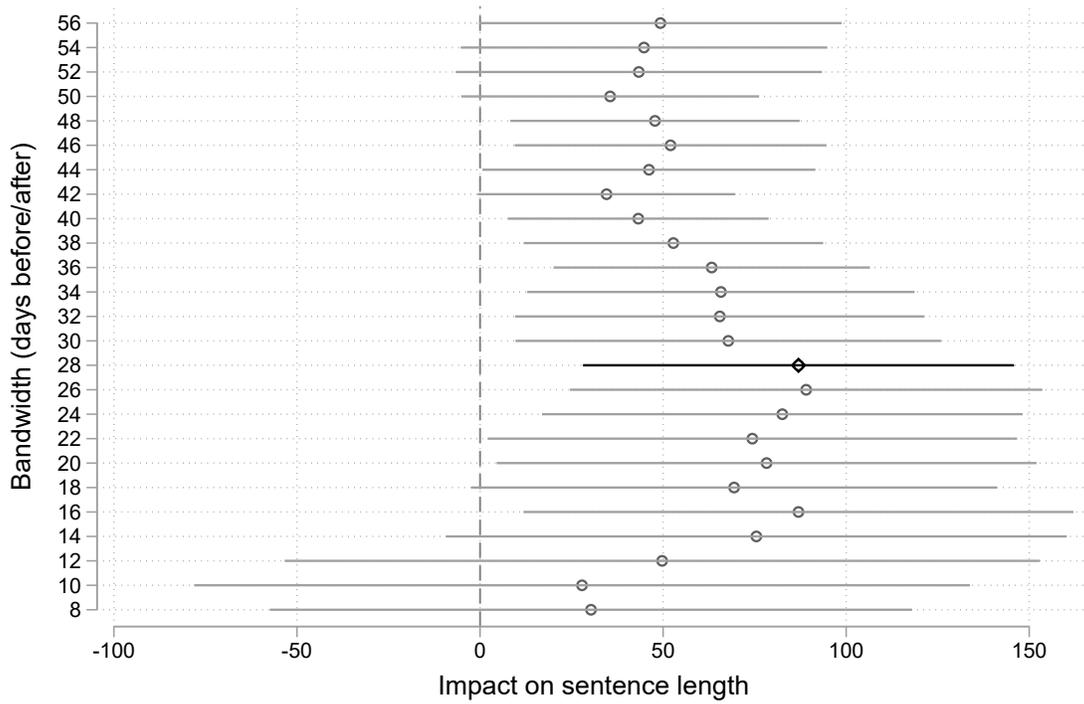
Notes: This figure plots the estimated coefficients β_3 and 95% confidence intervals for the regression corresponding to Equation 1 with sentence length (conditional on prison) as the dependent variable. We drop observations from our sample depending on the specification: excluding drugs crimes, excluding property crimes, excluding violent crimes, excluding traffic crimes, excluding other crimes. The dashed red vertical line marks the effect size estimated with the main specification (in Table 3). Standard errors are clustered at the judicial district-year level. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A8: Placebo impact of shock in salience on sentence length for suspects of Moroccan descent, around September 18 in previous years



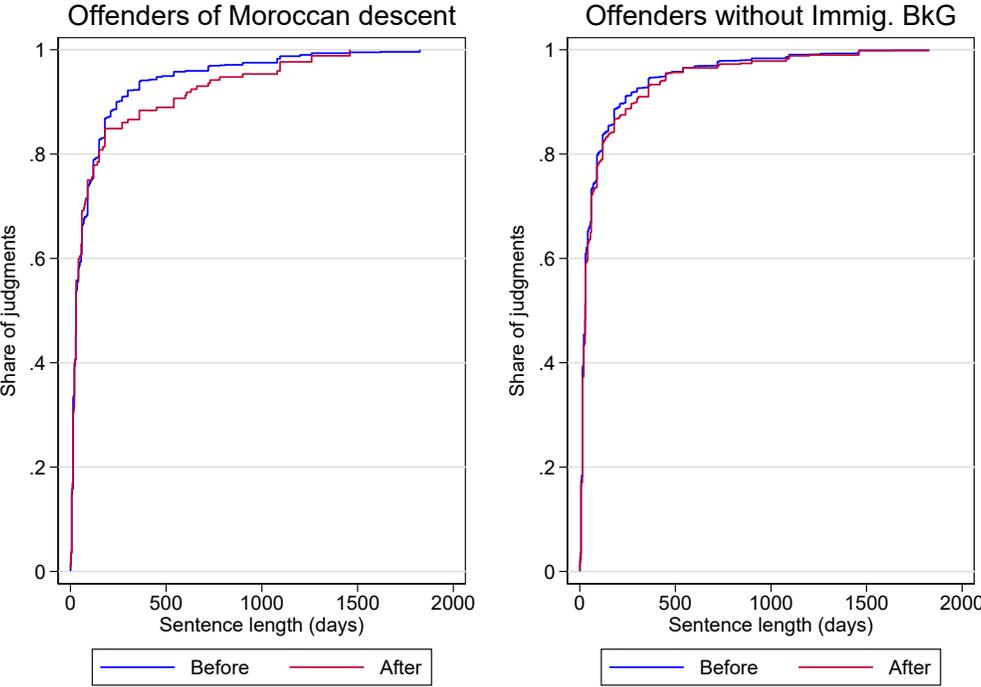
Notes: This figure plots the estimated coefficients β_3 and 95% confidence intervals for the regression corresponding to Equation (1) with sentence length as the dependent variable. We change the dummy variable *ShockYear* in each specification to 2015, 2016, 2017, 2018 as placebo shock years, and to 2019 as our “true” shock year. We use the sample conditional on prison sentence and depending on the year we use as *ShockYear* we change our sample to that year and the years before (e.g. if *ShockYear* = 2015 we use years 2014–2015). Standard errors are clustered at the judicial district-year level. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A9: Impact of shock in salience on sentence length for suspects of Moroccan descent, using different bandwidths around shock



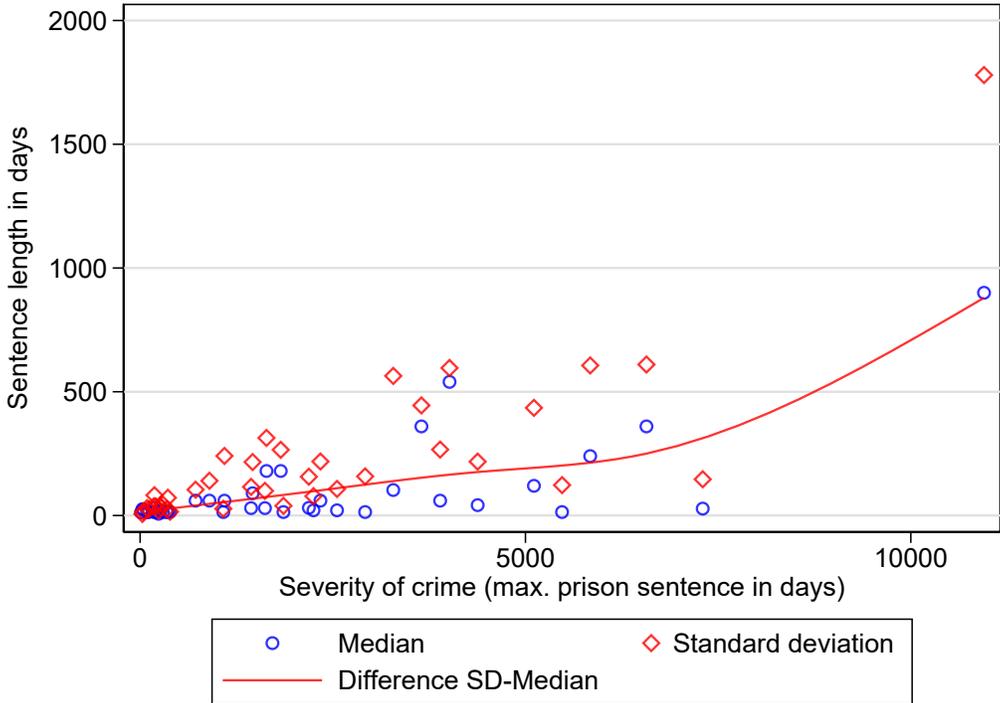
Notes: This figure plots the estimated coefficients β_3 and 95% confidence intervals for the regression corresponding to Equation (1) with sentence length as the dependent variable. We use the sample conditional on prison sentence and change the bandwidth (days before/after the shock) in each specification and change the restriction on lead time accordingly (e.g., if we use a bandwidth of 56 days, a case should have been sent by the public prosecutors 56 days before the shock). The estimate of our effect using the main specification is indicated in red. Standard errors are clustered at the judicial district-year level. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A10: Cumulative density distribution of sentence length for suspects of Moroccan descent and those without an immigration background, before and after September 18, 2019



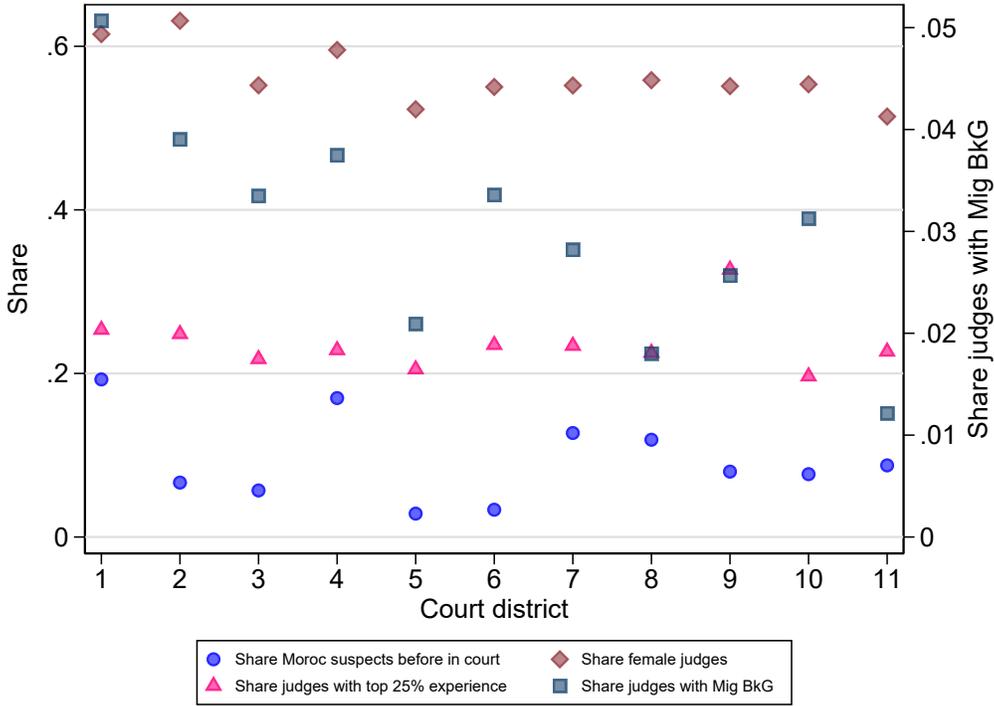
Notes: This figure plots the cumulative density of the given sentence length (conditional on incarceration) for suspects of Moroccan descent (left panel) and suspects without an immigration background (right panel), between August 21 and September 18, 2019 in blue and between September 18 and October 15, 2019 in red. The distribution is right-censored at 2000 days. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A11: Variation in sentencing: Median and standard deviation of actual sentence lengths, categorized by severity of crime (max. sentence length possible)



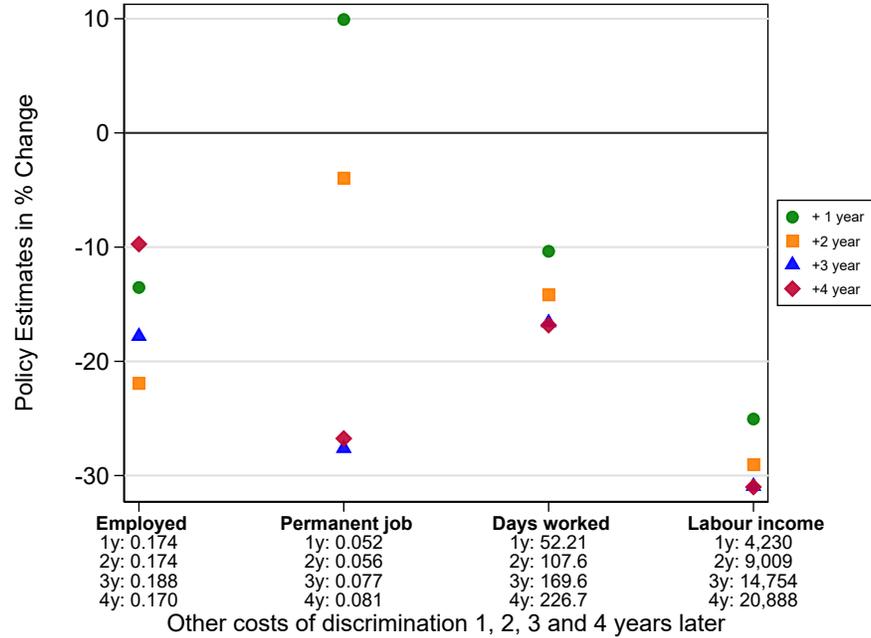
Notes: This figure plots the median (blue dots) and standard deviation (red squares) of actual sentence lengths given by judges (in months, y-axis), categorized by crime severity, measured as the crime’s maximum sentence length (x-axis). The red line represents a smoothed fit of the difference between the standard deviation and the median, illustrating trends in the variability of sentences across different crime severities. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A12: Mean court characteristics



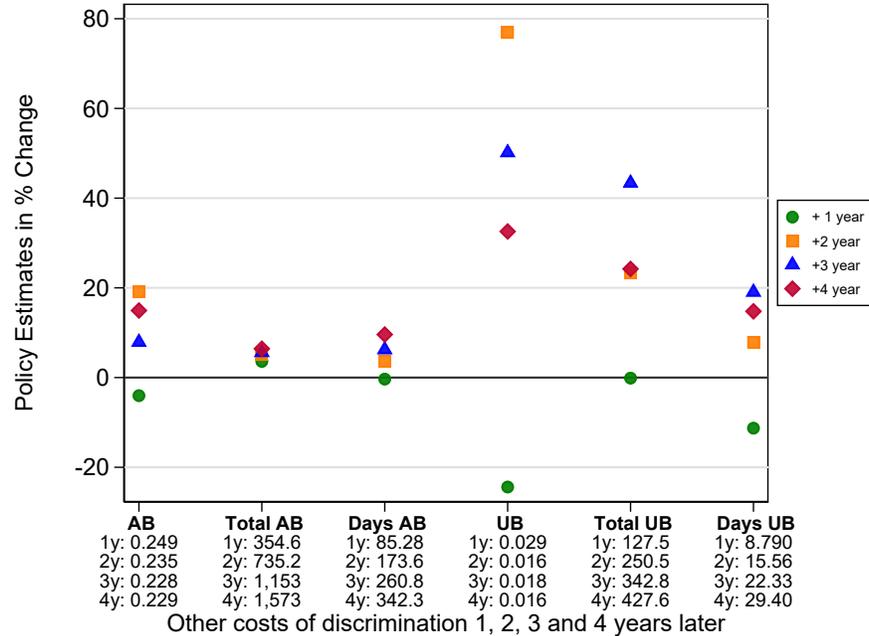
Notes: This figure plots the average court characteristics by judicial district for September 2019. Court characteristics are the share of previous cases with a defendant of Moroccan descent appearing before the court pre-shock (blue circle), share of female judges (red diamond), share of judges in top 25% national distribution of experience (pink triangle), share of judges with a non-western immigration background (blue square). Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands and publicly available data on judges at <https://namenlijst.rechtspraak.nl/>.

Figure A13: Impact of shock in salience on labor market outcomes for defendants of Moroccan descent



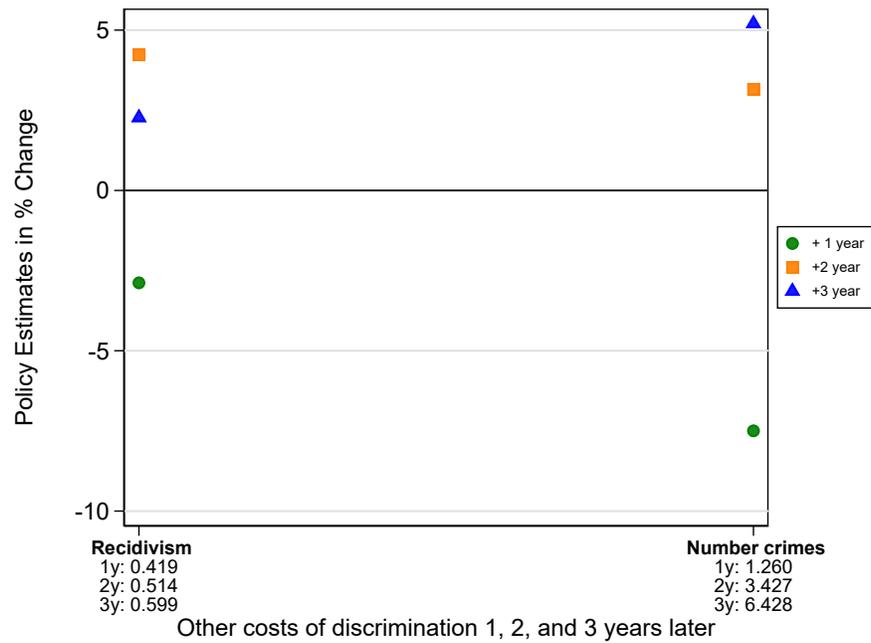
Notes: This figure plots the estimated coefficients from a simplified version of Equation 1 that does not use the third difference. We exclude previous years as control years to compare suspects who (possibly) experience similar labor market discrimination due to the salience shock but differ in the discrimination they faced in the criminal justice system. We look at labor market outcomes 1, 2, 3, and 4 years following judgment. Only about 3% (in the third year) and 1.5% (in the fourth year) of defendants in our sample remain incarcerated. The dependent variables are dummies for being employed employed, having a permanent contract, the (cumulative) number of days worked and the (cumulative) amount of labour income. Point estimates are transformed to differences in percentages relative to the baseline mean outcome for defendants of Moroccan descent (before the shock). Absolute coefficients and standard errors in brackets. Employed: 1y: -0.024 (0.034), 2y: -0.038 (0.036), 3y: -0.034 (0.031), 4y: -0.017 (0.023). Permanent job: 1y: 0.005 (0.017), 2y: -0.002 (0.017), 3y: -0.021 (0.021), 4y: -0.022 (0.022). Days worked: 1y: -5.411 (7.980), 2y: -15.278 (14.294), 3y: -28.183 (23.695), 4y: -38.152 (30.847). Labour income: 1y: -1059.544 (896.030), 2y: -2615.412 (1772.693), 3y: -4563.365 (3033.745), 4y: -6476.767 (4223.170). Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A14: Impact of shock in salience on benefit receipts for defendants of Moroccan descent



Notes: This figure plots the estimated coefficients from a simplified version of Equation 1 that does not use the third difference (previous years) with as dependent variables: a dummy variable that indicates whether a suspect is on welfare/assistance benefits (AB), the (cumulative) total amount of received AB, the (cumulative) number of days AB, a dummy variable that indicates whether a suspect receives unemployment benefits (UB), the (cumulative) total amount of received UB, the (cumulative) number of days UB 1, 2, 3 and 4 years after judgment. Point estimates are transformed to differences in percentages relative to the baseline mean outcome for defendants of Moroccan descent (before the shock). Absolute coefficients and standard errors in brackets. Assistance benefits (AB): 1y: -0.010 (0.035), 2y: 0.045 (0.037), 3y: 0.018 (0.020), 4y: 0.034 (0.029). Total AB: 1y: 12.759 (62.111), 2y: 37.868 (91.595), 3y: 64.613 (114.738), 4y: 101.101 (155.562). Days AB: 1y: -0.297 (10.915), 2y: 6.314 (21.166), 3y: 16.211 (25.794), 4y: 32.868 (31.547). Unemployment benefits (UB): 1y: -0.007 (0.007), 2y: 0.012 (0.012), 3y: 0.009 (0.010), 4y: 0.005 (0.007). Total UB: 1y: -0.157 (92.293), 2y: 58.481 (126.519), 3y: 148.628 (165.432), 4y: 103.532 (192.801). Days UB: 1y: -0.992 (2.719), 2y: 1.226 (3.986), 3y: 4.254 (6.093), 4y: 4.340 (7.341). Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Figure A15: Impact of shock in salience on recidivism for defendants of Moroccan descent



Notes: This figure plots the estimated coefficients from a simplified version of Equation 1 that does not use the third difference (previous years) with, as dependent variables: a dummy variable that indicates whether there is recidivism by the suspect and the (cumulative) number of crimes 1, 2, and 3 years after judgment (data for 4 years is unavailable). Point estimates are transformed to differences in percentages relative to the baseline mean outcome for defendants of Moroccan descent (before the shock). Absolute coefficients and standard errors in brackets. Recidivism: 1y: -0.012 (0.031), 2y: 0.022 (0.023), 3y: 0.014 (0.023). Number of crimes: 1y: -0.094 (0.175), 2y: 0.108 (0.479), 3y: 0.334 (0.805). Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

B Additional Tables

Table B1: Impact of shock in salience on sentence length for suspects of Moroccan descent, detailed table

	Court Decisions			
		Sentence Length		
	(1)	(2)	(3)	(4)
Moroccan descent	98.879**	94.505**	85.495***	86.980***
x Post	(42.592)	(38.955)	(29.878)	(30.024)
Time F.E.	No	Yes	Yes	Yes
District F.E.	No	Yes	Yes	Yes
Crime F.E.	No	No	Yes	Yes
All Individ. Controls	No	No	No	Yes
Mean before	122.6	122.6	122.6	122.6
Share treated	20%	20%	20%	20%
Observations	11,772	11,375	11,374	11,374

Notes: The table shows regression results corresponding to the estimate of coefficient β_3 in Equation (1), where fixed effects and control variables are added sequentially. We use the sample conditional on a court imposing a prison sentence (at first instance). The dependent variable is a continuous variable indicating the number of days of prison. Robust standard errors clustered at the judicial district-year level in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

Table B2: Balance table suspect and case characteristics around time of shock, conditional on prison sentence

	Court		
	non-imm. (1)	Moroc. (2)	diff-in-diff (3)
Suspect:			
Age at crime	37.64 (12.10)	30.17 (9.458)	1.186 (1.120)
Criminal hist.	3.021 (3.188)	3.767 (3.289)	0.225 (0.361)
Male	0.904 (0.295)	0.941 (0.235)	0.001 (0.033)
Partner	0.120 (0.325)	0.047 (0.211)	0.011 (0.022)
Children	0.452 (0.498)	0.225 (0.418)	0.021 (0.057)
Basic educ.	0.328 (0.470)	0.294 (0.456)	0.010 (0.046)
[Earnings>0]	0.185 (0.388)	0.121 (0.327)	- 0.0056 (0.014)
[Benefits>0]	0.615 (0.487)	0.579 (0.494)	- 0.005 (0.075)
Crime type:			
Property	0.424 (0.494)	0.445 (0.497)	0.058 (0.064)
Violent	0.225 (0.418)	0.159 (0.366)	- 0.019 (0.039)
Drugs	0.068 (0.251)	0.108 (0.310)	0.015 (0.030)
Traffic	0.155 (0.362)	0.169 (0.375)	- 0.075 (0.065)
Other	0.128 (0.334)	0.120 (0.325)	0.022 (0.042)
Lead time	24.28 (23.25)	24.85 (27.84)	- 4.255 (3.005)
Time F.E.			Yes
Share treated			20.3%
Observations	9,013	2,361	11,374

Notes: This table shows the sample means (standard deviations in parentheses) split by Moroccan/no immigration background of the suspect for the subsample for cases at the court conditional on incarceration. Column (3) reports the estimated coefficient β_3 from Equation 1 with $Y_{i,c,t}$ being the variables listed in the rows. Criminal history is defined separately at each stage: it counts the number of times the suspect appeared before the court in the five years before. All suspect characteristics are measured before the crime. Lead time defines the time in weeks between when the public prosecutor sent the case to court and when the decision was made. Robust standard errors clustered at the judicial district-year level in parentheses. $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands on decisions made between 21 August and 15 October between 2014-2019.

Table B3: Impact of shock in salience on sentence length for suspects of Moroccan descent, using alternative specifications

	Sentence length							
	Detail. crime type (1)	Cond. convict. (2)	Simple pre-post (3)	DiD w/o 2014-2018 (4)	Top 10% Max (5)	Court crime type FEs (6)	Robust SEs (7)	Distr. clust. SEs (8)
Moroccan desc. x Post	74.522*** (26.807)	32.875** (13.759)	62.283 (36.058)	64.449 (37.565)	0.084** (0.035)	75.447*** (24.891)	86.980*** (27.615)	86.980** (31.436)
Time F.E.	Yes	Yes	No	No	Yes	Yes	Yes	Yes
District F.E.	Yes	Yes	No	No	Yes	Yes	Yes	Yes
Crime F.E.	Yes	Yes	No	No	Yes	Yes	Yes	Yes
Indiv. Controls	Yes	Yes	No	No	Yes	Yes	Yes	Yes
Mean before	122.6	51.08	98.05	98.05	0.172	122.6	122.6	122.6
Share treated	20.0%	18.5%	100%	20.0%	20.0%	20.0%	20.0%	20.0%
Observations	11,374	32,699	382	1,746	11,374	11,374	11,374	11,374

Notes: The table shows regression results corresponding to the estimate of coefficient β_3 in Equation (1), with as dependent variable a continuous variable indicating the number of days of prison. In Col. (1) we use more detailed crime type fixed effects, in Col. (2) we use the sample conditional on conviction, in Col. (3) we use the sample of only suspects with a Moroccan migration background in 2019, in Col. (4) we do not use 2014-2018 as additional control years, in Col. (5) we use a different dependent variable (a dummy variable that indicates whether the sentence length was in the top 10% of the maximum sentence length within the crime type), in Col. (6) we use crime type fixed effects as determined by the courts (instead of public prosecutor), in Col. (7) we use robust standard errors, in Col. (8) we use robust standard errors clustered at the judicial district-year-month level. If not listed otherwise, we use the sample conditional on a court imposing a prison sentence (at first instance) and robust standard errors clustered at the judicial district-year-month level in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$. Source: Results based on calculations by the authors using micro-level data from Statistics Netherlands.

C Data Appendix

Administrative data from Statistics Netherlands

We use administrative data from Statistics Netherlands, which contains information on all individuals in the Netherlands.¹ We start with the registry of police decisions (*VERDEINDTAB*), public prosecutor decisions (*UITSTROOMOMTAB*) and court decisions (*UITSTROOMRECHTERTAB*).² Public prosecutor and court cases can be linked directly using a shared case identifier, but this case identifier is not available in the police data. Police decisions are available since 2005, public prosecutor and court decisions since 2001. Due to policy changes in the criminal justice system before 2014, our main analysis sample starts in 2014. We do use the years before to construct a criminal history for all suspects, measured by the number of cases in a stage for each suspect in the five years before (and a dummy variable that equals one if a suspect has a criminal history). We also compute the exposure before to suspects of Moroccan descent as the share of cases decided in a police/judicial district in the three years before the decision of interest. We only keep cases with suspects with a(n anonymised) social security number (suspects registered at the “Gemeentelijke Basis Administratie (GBA)”) in order to link the crime data to other administrative registries containing demographic and socioeconomic characteristics. We drop crimes committed by legal persons. We have information on the police district where a case was handled, but do not receive direct information on the court (public prosecutor’s office) in the court (public prosecutor) data. Instead, we have information on the municipality where a crime was committed. We aggregate these municipalities to judicial districts (“arrondissement”) and use that as a proxy for the actual judicial district where a case is handled.³ We classified crimes into five different crime types: property crimes (“scm10” code between 1000000-2000000),

¹The administrative data from Statistics Netherlands is available at a remote-access facility after signing a confidentiality agreement.

²The detailed police data used is not standardly available in catalogue of Statistics Netherlands. More information available upon request.

³According to the Dutch criminal procedural law (*Artikel 2 Wetboek van Strafvordering*) courts are authorized if the crime is committed in the judicial district; the suspect lives in the judicial district; the suspect is located in the judicial district; the suspect was last known to live in the judicial district; the prosecution in another case already started in that judicial district. In this paper we assume the first assignment rule is used.

violent and sexual crimes (“scm10” code 2020101, 2020102 or between 3000000 and 4000000), traffic offenses (“scm10” code between 5000000 and 6000000), drugs offenses (“scm10” code between 6000000 and 7000000), and other crimes (“scm10” code between 2000000-3000000, 4000000, 7000000 or larger than 9000000).

We link suspects to the registry of persons (*GBAPERSOONTAB*) to obtain demographic information like birth date (to compute age at crime), gender and immigration background. Immigration background includes not only the first, but also the second generation. Next, we use the parent-child registry (*KINDOUDERTAB*) to link the suspects in our sample to their parents so we can determine the immigration background of their parents. Hence, we are also able to identify which suspects have a third generation immigration background.⁴

For information on socioeconomic characteristics of suspects (measured in the year before the decision of interest) we use various registries. We use *KINDOUDERTAB* to construct variables on whether a suspect has children, how many, and if they have at least one young child (under age 4). We use *INPATAB* (tax registry) for information on the income percentile of a suspect, and whether a suspect earns income, is economically independent (min. 70 percent of minimum income), unemployed (receives unemployment benefit), ill (receives illness benefits), disabled (receives disability benefit), on welfare (receives “bijstand”), receives other types of social security, is the head of a household with a partner, is the head of a household without a partner, is a minor child in a household, is a non-minor child in a household, is a child/student without income in a household, is a child/student with income in a household. We use *HOOGSTEOPLTAB* to determine whether a suspect obtained more than secondary education (e.g. vocational post-secondary education, applied university, university) and *ONDERWIJSINSCHRTAB* on whether a suspect is registered in school. We use *GBAADRESOBJECTBUS* and *VSLGW TAB* to determine whether a suspect lived in a neighbourhood in Utrecht which has ties with the “Moroccan mafia”.⁵

For the longer term labour market outcomes (1 to 4 years after the decision of interest)

⁴If an individual has the Netherlands as country of origin (meaning they do not have a first or second generation immigration background), we look at the country of origin of their mother (if available). If the country of origin of their mother is also the Netherlands, we look at the country of origin of their father (if available).

⁵According to the media, these are the following larger neighbourhoods (“wijken”): Overvecht, Vianen bedrijventerrein, Vianen centrum, Vianen buitengebied, and the smaller neighbourhoods (“buurten”): Zuilen-Noord, Kanaleneiland-Noord and Kanaleneiland-Zuid.

we use slightly different registries that are available for more recent years than the preferred used registry before. We use *SPOLISBUS* to construct information on whether a suspect is employed, the labour income of a suspect, the number of days worked, whether they have a permanent contract and the number of days worked with a permanent contract. We use *SECMBIJSTMNDBEDRAGBUS* and *SECMWERKLMNDBEDRAGBUS* to construct variables on whether a suspect received assistance or unemployment benefits, the total amount and the total number of assistance or unemployment benefits. We use *VERDEIND* to construct variables on recidivism and the number of crimes committed after.

Appeal case data from the judiciary

We enrich the administrative data on the criminal justice system with additional data from the Judiciary on decisions on appeal between 2014-2023.⁶ We link this information to the court cases in the administrative data using case identifiers. We construct our outcome variable by updating sentence length as imposed by the court of first instance with the sentence length in appeal (if sentence length was changed in appeal). If no prison sentence is imposed in appeal, sentence length is set equal to zero.

News data from Nexis Uni

We collected the data on news articles from [Nexis Uni](#). We used the search terms (and/or): *Taghi, Marengo, Mocro maffia, Marokkaanse maffia, Marokkaanse criminaliteit, Marokkaanse crimineel, Marokkaanse criminelen*. We exclude news articles about the Dutch TV series called “Mocro Maffia”, by excluding news articles that also contain the following search terms: *serie, series, misdaadseries, misdaadserie, dramaserie, dramaseries*. We limited our search to news articles published between 2014–2021 in the main Dutch national newspapers: *AD, De Telegraaf, NRC, Nederlands Dagblad, Reformatorisch Dagblad, FD, Trouw, De Volkskrant*. We downloaded the metadata of all articles corresponding to these criteria and created a daily panel dataset with the number of articles and the number of words that contained our search terms.

⁶The non-public data can be requested through a [formal procedure](#) at “de Rechtspraak”.

Data on judge characteristics

The administrative court data from Statistics Netherlands does not contain judge identifiers or data on judges, so we collected data on judges from a public registry.⁷ We scraped data on the names, gender, and current and previous positions (job type, start and end date and judicial district) of judges.

The dataset does not contain information on immigration background of judges. We therefore determine immigration background manually based on surname. We asked a research assistant to assign immigration background, with the options: no, non-western, or western immigration background.⁸ One of the authors did the same independently.⁹ We agreed on 97.35% of the judges classified as having a non-western immigration background. For the 122 judges with a conflicting coding, we used the international name database “Forebear” and coded judges as having a non-western immigration background if the most common country of origin of their name was from Africa, Latin-America, Asia or Turkey. For judges with two names (all of them women), we assume the second name is the maiden name and therefore use the coding of the second name.

We compute months of experience based on time since first position at a court. Next we create a dummy variable that equals one if the years of experience of a judge is in the top 25 percent of the distribution of all judges in that month and year. The court data from Statistics Netherlands does not allow us to link cases to judge-level data. We therefore collapse the judge data to a monthly dataset at the judicial district level, leaving us with the share of judges with an immigration background, the share of female judges, and the share of judges in the top 25 percent distribution of experience nationally, and link this dataset to

⁷Data on judges can be found at <https://namenlijst.rechtspraak.nl/>. This public registry is not easy to access and was therefore made more accessible on the website [Open Rechtspraak](#) by the Open State Foundation. We used their website to scrape judge characteristics on September 17, 2023.

⁸The divide between non-western and western immigration backgrounds used to be made by Statistics Netherlands. They defined an immigration background as non-western if the country of origin is Africa, Latin-America, Asia or Turkey.

⁹Surinamese names are often difficult to distinguish from Dutch names. We therefore checked whether judge names appeared in a [publicly available list](#) with Surinamese names. If names appeared in this list, we checked in [the Dutch name database](#) (“CBG”) if those names are only Surinamese or also/primarily Dutch. We code names as Dutch if the description of the name only mentions a Dutch history; both if the description mentions both a Dutch and Surinamese history; and Surinamese if the description mostly/only mentions a Surinamese history; unknown if no description available. We code the names classified as both or unknown as Surinamese if a name occurred less than 500 times (in 2007).

court cases based on district and decision month.

Text analysis of publicly available court rulings

We collected all publicly available court rulings (by courts in first instance) in criminal cases between 21 August 2019 and 15 October 2019, published at the website of the Judiciary.¹⁰ The number of published cases is 593, while we observe 11,551 cases in the same time period in our court data from Statistics Netherlands. The published sample of cases is not random: criminal cases at courts in first instance are i.a. published if they concern a crime against life, or when a prison sentence of at least four years or “TBS” (custodial measure for mentally-ill offenders) is imposed, or when a case has received media attention. However, this list is not exhaustive and the Judiciary has expressed the aim to publish cases that do not only contain the standard wordings as much as possible.¹¹ Another important limitation is that published rulings are pseudonymized, making it impossible to identify the offender and their immigration background. Hence, we can only look at a change in case characteristics of all cases after the shock in salience on 18 September 2019, without distinguishing cases with an offender of Moroccan descent from cases with an offender without an immigration background.

Next, we analyzed the text of the rulings to collect more information on behavior of defendants and their lawyers in courts. Our research assistant and one of the authors both randomly picked and read 30 court rulings in the first/last week of our time period of interest to determine what types of behavior could be identified. We agreed on the following information of interest: a ruling mentioned Morocco (e.g. if suspect is born there), defendant/lawyer requests acquittal, defendant cooperates, defendant confesses, defendant denies, recusal of court, defendant appears at hearing. Per court ruling, binary variables were created that equal one if a ruling contained search terms, listed at the end of this section, related to these topics. Additionally, our research assistant coded the number of words spent by the defendant on evidence (“bewijs standpunt verdediging”) and the number of words spent by the defendant on the punishment (“straf standpunt verdediging”) manually.

¹⁰<https://uitspraken.rechtspraak.nl/>

¹¹The criteria for publication of court rulings can be found at [the website of the Judiciary](#).

Table C1 shows the difference in case characteristics in the four weeks after the shock in salience compared to before. No significant changes are found that could indicate that the behavior of offenders or their lawyers changed for the worse. The only thing that changed significantly is that offenders are more likely to request acquittal after 18 September 2019 than before. If anything, that suggests that offenders and/or their lawyers try to exert more effort (i.e. a positive change in behavior) after the shock than before. A more general specification of acquittal (see list below for search terms) does not show a similar significant change.

Table C1: Difference in case characteristics before and after shock in salience

	Court judgment contains									
	Morocco	Acquittal	Acquittal general	Cooperation	Confess	Denial	Non- appearance	Recusal	Evidence word count	Punishment word count
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Post	0.003 (0.017)	0.119*** (.035)	0.040 (.052)	-0.068 (0.043)	-0.002 (.071)	0.0423 (.035)	-0.011 (.024)	-0.007 (0.007)	-18.140 (11.19)	-7.022 (7.870)
Time F.E.	No	No	No	No	No	No	No	No	No	No
District F.E.	No	No	No	No	No	No	No	No	No	No
Mean before	0.017	0.229	0.804	0.338	0.783	0.133	0.088	0.013	161.8	71.47
Observations	593	593	593	593	593	593	593	593	486	360

Notes: The table shows the estimates of coefficient β_2 in $Y_{c,t} = \beta_1 + \beta_2 Post_t + \epsilon_p$. We use the sample of cases judged between August 21 and October 15 in 2019. Robust standard errors clustered at the judicial district level are in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$. Source: Results based on calculations by the authors using published court cases on [the website of the Judiciary](#).

Search terms Morocco: “*marokko*”. Acquittal: “*bepleit een algehele vrijspraak*”, “*bepleit algehele vrijspraak*”, “*bepleit vrijspraak*”, “*heeft vrijspraak bepleit*”, “*bepleit verdachte vrij*”, “*standpunt gesteld dat niet wettig en overtuigend bewezen*”, “*bepleit is dat de verdachte dient te worden vrijgesproken*”, “*heeft betoogd dat verdachte moet worden vrijgesproken*”, “*heeft integrale vrijspraak bepleit*”, “*bepleit integrale vrijspraak*”, “*verdachte dient te worden vrijgesproken*”, “*verdachte moet worden vrijgesproken*”. More general acquittal (which increases the likelihood that we are able to identify a request for acquittal, but is also more prone to false positives): “*vrijspraak*”, “*vrijgesproken*”. Cooperation (using combinations of search terms that should appear within 500 characters from each other): “*verdachte*”

“mee” “werken”, “verdachte” “werkt” ‘mee”, “verdachte” ”problemen” “werken”, “meewerken” “verdachte”. Confession: “bekennende verdachte”, “verdachte” “bekent” (within 500 characters), “bewezenverklaring”. Denial: “ontkent”, “ontkennen”. Appears at hearing: “verstek”, “niet verschenen”. Recusal: “waking”.