

# Uncertainty and Economic Activity: A Global Perspective

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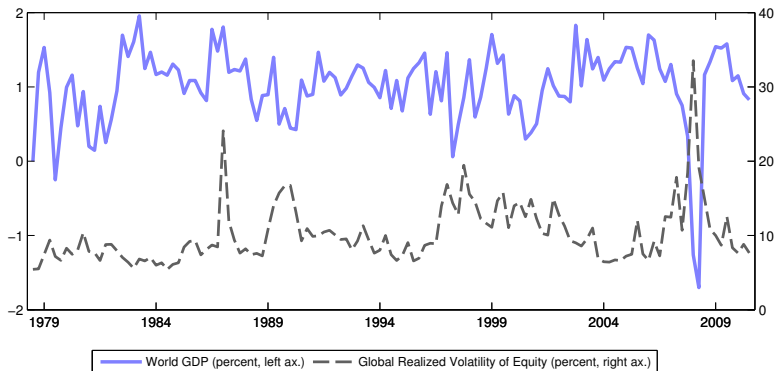
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<sup>1</sup>Bank of England and CfM. The views expressed in this paper are solely those of the authors and should not be taken to represent those of the Bank of England.

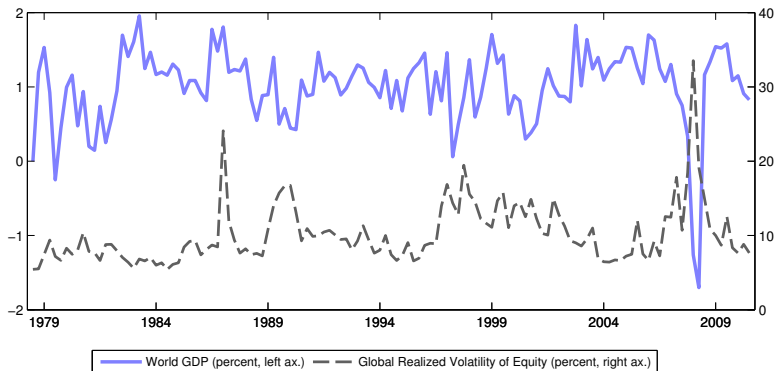
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# Strong correlation between “uncertainty” and economic activity



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- ▶ During the crisis increase in uncertainty/volatility and contraction in activity
- ▶ After the crisis, low volatility and a recovery of economic activity

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  - Compute quarterly country-specific realized volatility measures (as a proxy for economic uncertainty) using daily returns of 109 asset prices worldwide
  - Set up a factor model for volatility and the business cycle in which both are driven by the same set of global common factors
  - Exploit the different cross-country correlation structure of volatility and GDP growth to identify the factors and the shocks

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  - Exploit the different cross-country correlation structure of volatility and GDP growth to identify the factors and the shocks
- ▶ **What do we find?** Show that conditional on global factors there is little correlation left between volatility and economic activity

# Outline

1. A simple factor model of volatility and macro dynamics
2. Results

# A standard model of volatility and economic activity

- ▶ Model used in the literature (abstracting from dynamics) to interpret correlation between  $v_t$  and  $\Delta y_t$

$$v_t = \alpha \Delta y_t + \varepsilon_t$$

$$\Delta y_t = \beta v_t + u_t$$



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- ▶ Since the covariance matrix  $Cov(v_t, \Delta y_t)$  provides only three moments, the system is not identified
- ▶ Identification of structural parameters and shocks is typically achieved with an exclusion restriction (ie  $\alpha = 0$  or  $\beta = 0$ )

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$$v_t = \lambda n_t + \varepsilon_t$$
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$$\begin{aligned}v_t &= \lambda n_t + \varepsilon_t \\ \Delta y_t &= \gamma n_t + u_t\end{aligned}$$

- ▶ Again, identification of structural parameters and shocks cannot be achieved unless we impose an exclusion restriction, ie by assuming  $\lambda = 0$  (or  $\gamma = 0$ )

# An alternative model based on common factors: multi-country perspective

- ▶ Replace model above with the following disaggregated system of equations:

$$\begin{aligned}v_{it} &= \lambda_i n_t + \varepsilon_{it} && \text{for } i = 1, 2, \dots, N \\ \Delta y_{it} &= \gamma_i n_t + u_{it} && \text{for } i = 1, 2, \dots, N\end{aligned}$$

- ▶ Global volatility ( $v_t$ ) and world GDP growth ( $\Delta y_t$ ) are aggregates over a large number of countries

$$v_t = \sum_{i=1}^N \hat{w}_i v_{it}, \quad \Delta y_t = \sum_{i=1}^N w_i \Delta y_{it},$$

# Identifying assumptions

- ▶ **Assumption 1** Weights  $\mathbf{w} = (w_1, w_2, \dots, w_N)'$  are of order  $1/N$

$$\|\mathbf{w}\| = O_p(N^{-\frac{1}{2}}), \quad \frac{w_i}{\|\mathbf{w}\|} = O_p(N^{-\frac{1}{2}}) \quad \forall i,$$

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- ▶ **Assumption 2** The volatility innovations ( $\varepsilon_{it}$ ) are strongly correlated across countries, whilst the output innovations ( $u_{it}$ ) are weakly cross-correlated across countries.

$$\lambda_{\max}(\boldsymbol{\Sigma}_\varepsilon) = O_p(N) \quad \text{and} \quad \lambda_{\max}(\boldsymbol{\Sigma}_u) = O_p(1)$$

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- One cross-sectional unit plays dominant role in global financial markets but not in world activity
- ▶ These assumptions are not contradicted by the time series evidence



# Identification of the factor by aggregation

- ▶ Consider the cross-country weighted averages of the disaggregated system

$$v_t = \lambda n_t + \bar{\varepsilon}_t$$

$$\Delta y_t = \gamma n_t + \bar{u}_t$$

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- ▶ For  $N$  sufficiently large  $n_t$  can be identified from macro equation

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$$n_t = \gamma^{-1} \Delta y_t + \underbrace{\bar{u}_t}_{O_p(N^{-1/2})}$$

- ▶ However, since the volatility shocks ( $\varepsilon_{it}$ ) are strongly correlated across countries,  $n_t$  cannot be approximated by  $v_t$

# Implications: volatility equation

- ▶ Substitute  $n_t$  into the volatility equation to get

$$\begin{aligned} v_{it} &= \lambda_i \underbrace{\gamma^{-1} \Delta y_t + \bar{u}_t}_{n_t} + \varepsilon_{it} = \\ &= (\lambda_i \gamma^{-1}) \Delta y_t + \varepsilon_{it} + O_p(N^{-1/2}) \end{aligned}$$

- ▶ **Result** OLS consistently estimate the impact of contemporaneous changes in global activity on country-specific volatility
- ▶ **Result** We can identify the volatility innovation  $\varepsilon_{it}$

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- ▶ Get volatility innovations  $\hat{\varepsilon}_{it}^{eq}$
- ▶ Check validity of identifying assumption: strong cross-sectional dependence of volatility innovations  $\hat{\varepsilon}_{it}^{eq}$

# Pairwise correlation of volatility innovations: strong cross-sectional dependence

Table : PAIRWISE CORRELATION OF THE VOLATILITY INNOVATIONS

|         | PC   |       | PC   |        | PC   |
|---------|------|-------|------|--------|------|
| ARG     | 0.25 | INDIA | 0.30 | PHLP   | 0.35 |
| AUSTLIA | 0.53 | INDNS | 0.30 | SAFRC  | 0.51 |
| AUSTRIA | 0.44 | ITALY | 0.44 | SARBIA | 0.33 |
| BEL     | 0.49 | JAPAN | 0.47 | SING   | 0.52 |
| BRA     | 0.13 | KOR   | 0.40 | SPAIN  | 0.53 |
| CAN     | 0.56 | MAL   | 0.38 | SWE    | 0.54 |
| CHINA   | 0.47 | MEX   | 0.50 | SWITZ  | 0.55 |
| CHL     | 0.40 | NETH  | 0.54 | THAI   | 0.41 |
| FIN     | 0.21 | NOR   | 0.56 | TURK   | 0.35 |
| FRANCE  | 0.54 | NZLD  | 0.42 | UK     | 0.55 |
| GERM    | 0.56 | PER   | 0.49 | US     | 0.56 |

# Macro equation estimation

- ▶ Model country-specific GDP growth ( $\Delta y_{it}$ ) as:

$$\Delta y_{it} = \sum_{k=1}^p \Phi_{ik} \Delta y_{i,t-k} + \sum_{k=0}^q \Lambda_{ik} \Delta y_{t-k} + u_{it}$$

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- ▶ Get macro innovations  $\hat{u}_{it}$
- ▶ Check validity of identifying assumption: weak cross-sectional dependence of macro innovations  $\hat{u}_{it}$

# Pairwise correlation of GDP innovations: weak cross-sectional dependence

Table : PAIRWISE CORRELATION OF THE GDP INNOVATIONS

|         | PC    |       | PC    |        | PC    |
|---------|-------|-------|-------|--------|-------|
| ARG     | 0.00  | INDIA | 0.00  | PHLP   | 0.01  |
| AUSTLIA | 0.02  | INDNS | -0.02 | SAFRC  | 0.03  |
| AUSTRIA | -0.01 | ITALY | 0.03  | SARBIA | -0.02 |
| BEL     | -0.01 | JAPAN | -0.01 | SING   | -0.03 |
| BRA     | 0.02  | KOR   | 0.01  | SPAIN  | 0.01  |
| CAN     | 0.01  | MAL   | 0.01  | SWE    | 0.02  |
| CHINA   | -0.08 | MEX   | 0.02  | SWITZ  | 0.00  |
| CHL     | 0.01  | NETH  | 0.00  | THAI   | 0.03  |
| FIN     | 0.00  | NOR   | -0.01 | TURK   | -0.01 |
| FRANCE  | 0.01  | NZLD  | 0.04  | UK     | 0.01  |
| GERM    | -0.06 | PER   | 0.00  | US     | -0.04 |

**Note.** Lag length determined with Akaike criterion with a max of 4 lags.

# The impact of volatility innovations on economic activity

- ▶ Is there any relation left (after controlling for the global factor) between volatility and economic activity?

# The impact of volatility innovations on economic activity

- ▶ Is there any relation left (after controlling for the global factor) between volatility and economic activity?
- ▶ If yes, we should observe a (negative) correlation between  $\hat{u}_{it}$  and  $\hat{\varepsilon}_{it}^{eq}$
- ▶ Estimate the following country-specific equations

$$\hat{u}_{it} = b_i \hat{\varepsilon}_{it}^{eq} + \eta_{it} \quad \text{for } i = 0, 1, \dots, N$$



# The impact of volatility innovations on economic activity

Table : THE RELATION BETWEEN GDP INNOVATIONS AND VOLATILITY INNOVATIONS

|        | ARG   | AUSTLIA | AUSTRIA | BEL   | BRA   | CAN   | CHINA | CHL   | FIN   | FRANCE | GERM  |
|--------|-------|---------|---------|-------|-------|-------|-------|-------|-------|--------|-------|
| beta   | -0.04 | 0.00    | -0.01   | -0.02 | -0.02 | 0.00  | -0.01 | 0.01  | 0.00  | 0.01   | 0.01  |
| t-stat | -2.79 | -0.22   | -0.46   | -1.09 | -2.01 | -0.12 | -0.35 | 0.24  | 0.22  | 0.88   | 0.58  |
| R2     | 0.08  | 0.00    | 0.00    | 0.01  | 0.04  | 0.00  | -0.03 | 0.00  | 0.00  | 0.01   | 0.00  |
|        | INDIA | INDNS   | ITALY   | JAPAN | KOR   | MAL   | MEX   | NETH  | NOR   | NZLD   | PER   |
| beta   | -0.01 | -0.02   | 0.00    | -0.04 | -0.01 | -0.05 | 0.02  | 0.00  | -0.02 | -0.04  | 0.03  |
| t-stat | -0.61 | -1.27   | 0.20    | -2.05 | -0.61 | -2.83 | 0.82  | 0.19  | -1.06 | -1.75  | 0.74  |
| R2     | 0.00  | 0.02    | 0.00    | 0.03  | 0.00  | 0.06  | 0.00  | 0.00  | 0.01  | 0.03   | -0.04 |
|        | PHLP  | SAFRC   | SARBIA  | SING  | SPAIN | SWE   | SWITZ | THAI  | TURK  | UK     | US    |
| beta   | -0.02 | -0.01   | -0.02   | 0.02  | 0.00  | -0.02 | -0.01 | -0.01 | 0.02  | -0.01  | -0.02 |
| t-stat | -0.93 | -0.77   | -0.66   | 0.68  | 0.41  | -1.06 | -0.81 | -0.44 | 0.59  | -0.57  | -1.67 |
| R2     | 0.01  | 0.00    | 0.00    | 0.00  | 0.00  | 0.01  | 0.00  | 0.00  | 0.00  | 0.00   | 0.02  |

# Conclusions

- ▶ Quantify the relation between uncertainty and economic activity using a novel multi-country approach
- ▶ Exploit the different cross-country correlation structure of volatility and GDP growth to establish the direction of causation
- ▶ Much of the reduced form correlation between volatility and economic is driven by common (first moment) factors