Specification Testing with Prediction Criterion: Causality, Prediction, and External Validity

Introduction

Model specification testing.

- Regression analysis relies on the correctness of model specification.
- e.g., Durbin-Wu-Hausman test
- Correct model: orthogonality of the dependent variables and error term.

Prediction-based model specification test

- Assume availability of train data $\{(X_i, Y_i)\}_{i=1}^n$ and test data $\{\tilde{X}_j\}_{j=1}^m$.
- Test data: the data that we want to predict the outcome.
- There are only covariates \tilde{X}_{i} , and the target variables are unobservable.

New definition of correct models.

- Idea: If the model can predict target variables well, the model is correct.
- Under the definition, we show
- The asymptotic distribution of the least squares under covariate shift.
- The asymptotic distribution of the test statistics.

1. Covariate Shift Problem

Data-generating process (DGP):

There are two stratified data:

$$(X_i, Y_i) \sim p(x, y), \qquad (\tilde{X}_j, \tilde{Y}_j) \sim q(x, y),$$

where $X_i, X_j \in \mathbb{R}^d$ and $Y_i, \tilde{Y}_j \in \mathbb{R}$. \tilde{Y}_j is unobservable.

Observations:

$$\{(X_i, Y_i)\}_{i=1}^n \sim p(x, y), \qquad \{\tilde{X}_j\}_{j=1}^m \sim q(x),$$

Furthermore, we put the following assumption on the conditional pdf:

$$p(x, y) = p(y|x)p(x),$$
$$q(x, y) = p(y|x)q(x).$$

- p(y|x) is invariant across the two data.
- p(x) and q(x) can be changed
- p(x) and q(x) have a common support.

This setting is called learning under covariate shift.

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2. Definition of Correct Model

Linear model:

- Assume a linear model of $\mathbb{E}[Y_i|X_i]$ as $\mathbf{Z}^{\top}(\mathbf{X}_i)\boldsymbol{\beta}^*$.
- $Z(\cdot)$ is a mapping from X_i to some linear models.

Definition of correct model

- Our model specification is defined from the viewpoint of prediction.
- Parameter that minimizes the MSE over p(x, y) is defined as

$$\alpha_0 = \operatorname{argmin}_b \mathbb{E}_{p(x,y)} [(Y_i - Z^{\mathsf{T}}(X_i)b)^2].$$

Parameter that minimizes the MSE over q(x, y) is defined as

$$\gamma_0 = \operatorname{argmin}_b \mathbb{E}_{q(x,y)} \left[\left(\tilde{Y}_j - Z^{\mathsf{T}}(\tilde{X}_j) b \right)^2 \right].$$

- If $\alpha_0 = \gamma_0$, the model is specified correctly
- If $\alpha_0 \neq \gamma_0$, the model is misspecified.
- By using this definition, consider the following hypothesis:

 $\mathcal{H}_0: \alpha_0 = \gamma_0 \text{ and } \mathcal{H}_1: \alpha_0 \neq \gamma_0$

If \mathcal{H}_0 is rejected, the model specification is incorrect.

3. Covariate Shift Adaptation

However, we cannot observe \tilde{Y}_i .

Let us define a parameter estimated from $\{(X_i, Y_i)\}_{i=1}^n$ as

$$\widehat{\alpha} = argmin_b \widehat{\mathbb{E}}_{p(x,y)} \big[(Y_i - Z^{\top}(X_i)b)^2 \big],$$

where $\widehat{\mathbb{E}}_{p(x,y)}$ denotes the sample average of the samples from p(x,y).

Then, for $\{\tilde{X}_j\}_{j=1}^m$, we define the following estimator:

$$\widehat{\gamma} = argmin_b \widehat{\mathbb{E}}_{q(x,y)} \left[\left(\widetilde{Y}_j - Z^{\top}(\widetilde{X}_j) b \right)^2 \right]$$
$$\approx argmin_b \widehat{\mathbb{E}}_{p(x,y)} \left[(Y_i - Z^{\top}(X_i) b)^2 \frac{q(X_i)}{p(X_i)} \right]$$

• Thus, we approximate $\mathbb{E}_{q(x,y)}$ by using $\widehat{\mathbb{E}}_{p(x,y)}$ and $\frac{q(X_i)}{p(X_i)}$.

Let us denote the density ratio $\frac{q(x)}{p(x)}$ by $r^*(x)$.

We can estimate the density ratio with machine learning methods. e.g., uLSIF (Kanamori et al. (2012)).

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4. Double/Debiased Least Squares Estimator

Consider the asymptotic distribution of $\widehat{\gamma}$.

- The density ratio is estimated by machine learning methods.
- \rightarrow The estimator does not satisfy Donsker's condition.
- We use **double/debiased machine learning** to avid this problem.
- An estimator $\hat{\gamma}$ with a doubly robust form.
- Cross-fitting.
- Doubly robust estimator of the MSE over q(x, y):

$$\widehat{\mathbb{E}}_{q(x,y)}\left[\left(\widetilde{Y}_{j}-Z^{\mathsf{T}}(X_{j})b\right)^{2}\right]$$

$$\approx \widehat{\mathbb{E}}_{p(x,y)} \left[\left((Y_i - Z^{\mathsf{T}}(X_i)b)^2 - \left(\widehat{f}(X_i) - Z^{\mathsf{T}}(X_i)b \right) \right) \widehat{r}(X_i) \right] + \widehat{\mathbb{E}}_{q(x)} \left[\left(\widehat{f}(\tilde{X}_j) - Z^{\mathsf{T}}(\tilde{X}_j)b \right)^2 \right]$$

- $\hat{f}(x)$ is some consistent estimator of $f^*(x) = \mathbb{E}[Y_i \mid x]$.
- $\hat{r}(x)$ is some consistent estimator of $r^*(x) = \frac{q(x)}{p(x)}$.
- We construct the empirical MSE by using cross-fitting.
- Then, if n = m = N,

$$\sqrt{N}(\widehat{\boldsymbol{\gamma}}-\boldsymbol{\gamma}^*)=\boldsymbol{\mathcal{N}}(\boldsymbol{0},\widetilde{\boldsymbol{\Sigma}})$$

5. Hypothesis Testing

-We construct the test statistics to investigate the hypothesis

$$\mathcal{H}_0: \alpha_0 = \gamma_0 \text{ and } \mathcal{H}_1: \alpha_0 \neq \gamma_0$$

A standard choice is to use Wald statistics.

We can construct Wald statistics by using the estimators $\hat{\alpha}$ and $\hat{\gamma}$.

- **The Wald statistics follows** $\chi^2(k)$ distribution.
- k is the dimension of the linear model.
- We conduct hypothesis testing using the test statistics.
- If the null hypothesis reject, we can say that model is misspecified.

References

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